

Piecewise-linear Modeling of Multivariate Geometric Extremes

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Limit sets and gauge functions

- ▶ $\mathbf{X}_1, \dots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} f_{\mathbf{X}}$, d -dimensional, Exponential(1) margins...

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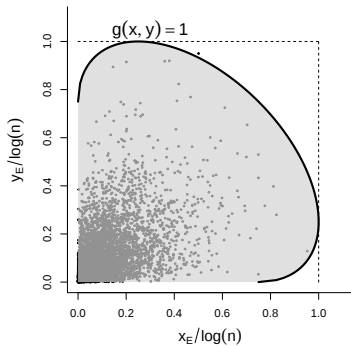
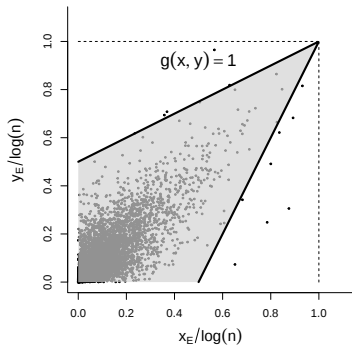
where the **gauge function**, $g : \mathbb{R}^d \longrightarrow \mathbb{R}_+$, is 1-homogeneous and continuous...

- ▶ ...then scaled sample clouds $\left\{ \frac{\mathbf{X}_1}{\log n}, \dots, \frac{\mathbf{X}_n}{\log n} \right\}$ converge onto a **limit set**,

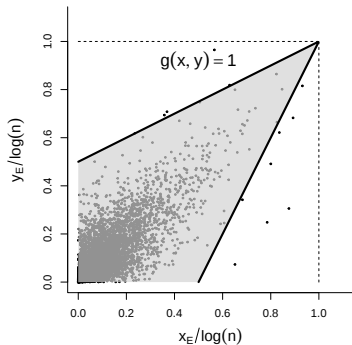
$$G := \left\{ \mathbf{x} \in \mathbb{R}^d \mid g(\mathbf{x}) \leq 1 \right\}$$

as $n \rightarrow \infty$ (Balkema and Nolde, 2010; Nolde and Wadsworth, 2022).

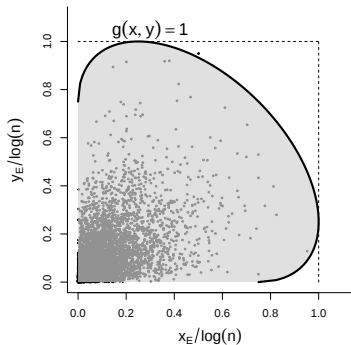
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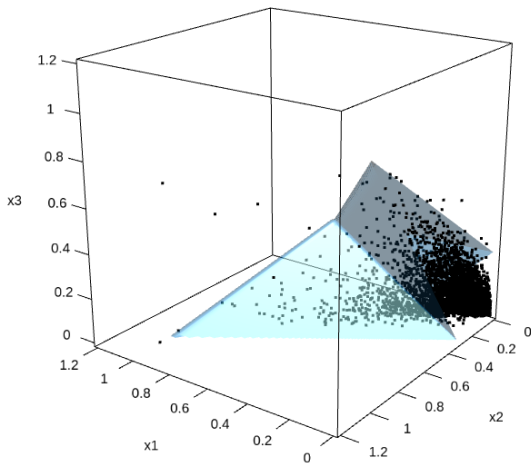
“pointy” $\implies$ AD



“not pointy” $\implies$ AI

Limit sets and gauge functions

Limit sets and gauge functions



Is this approach useful for extremal statistical inference across the multivariate tail?

If so, how can we estimate g from data and use it for extremal statistical inference?



Extremal inference with gauge functions

- ▶ Define $(R, \mathbf{W}) = (\|\mathbf{X}\|_1, \mathbf{X}/\|\mathbf{X}\|_1)$
- ▶ $f_{\mathbf{X}}$ satisfies $-\log f_{\mathbf{X}}(r\mathbf{w}) \sim rg(\mathbf{w})$ as $r \rightarrow \infty$.



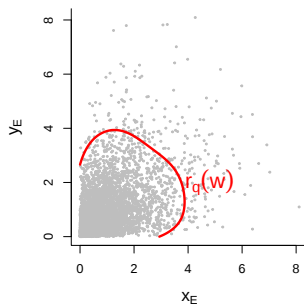
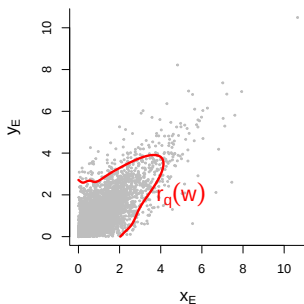
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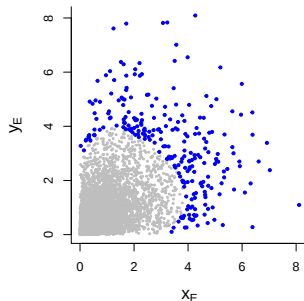
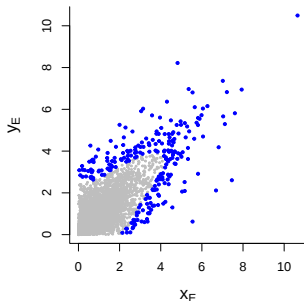
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Geometric approach: model fitting

- ▶ Wadsworth and Campbell (2024): fit the model

$$R \mid \{\mathbf{W} = \mathbf{w}, R > r_q(\mathbf{w})\} \sim \text{truncGamma}(\alpha, g(\mathbf{w}; \theta))$$

by maximizing

$$L(\theta; r_{1:n}, \mathbf{w}_{1:n}) = \prod_{i: r_i > r_q(\mathbf{w}_i)} \frac{f_{\text{Gamma}}(r_i; \alpha, g(\mathbf{w}_i; \theta))}{\bar{F}_{\text{Gamma}}(r_q(\mathbf{w}_i); \alpha, g(\mathbf{w}_i; \theta))}$$



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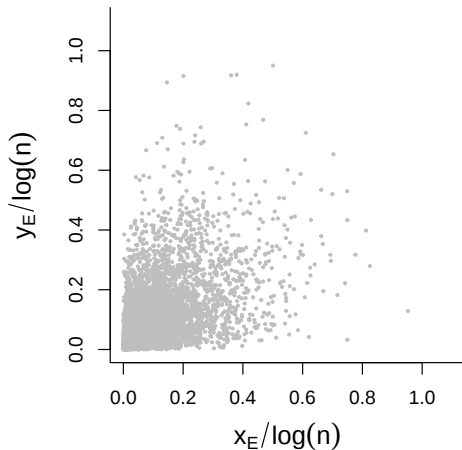
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- ▶ **Campbell and Wadsworth (2024): Define $g(\mathbf{w}; \theta)$ piecewise-linearly.**

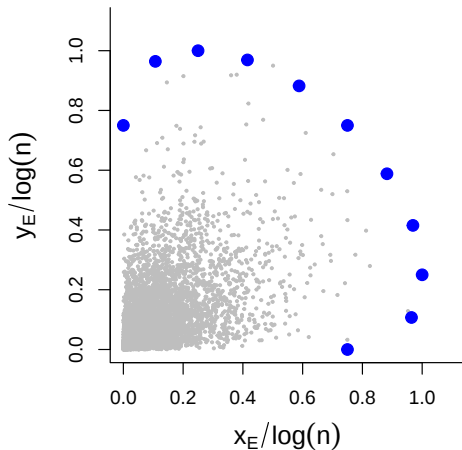


Semiparametric piecewise-linear approach



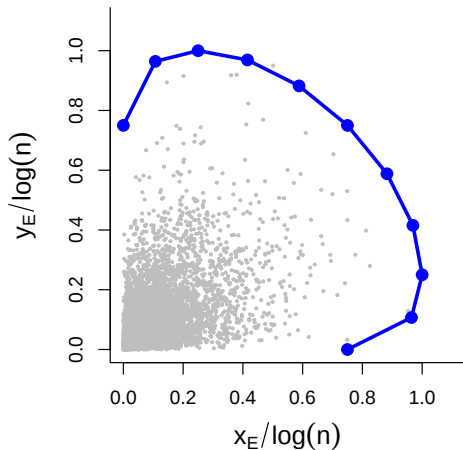


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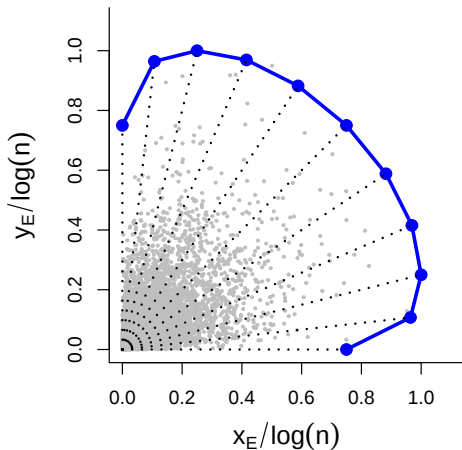


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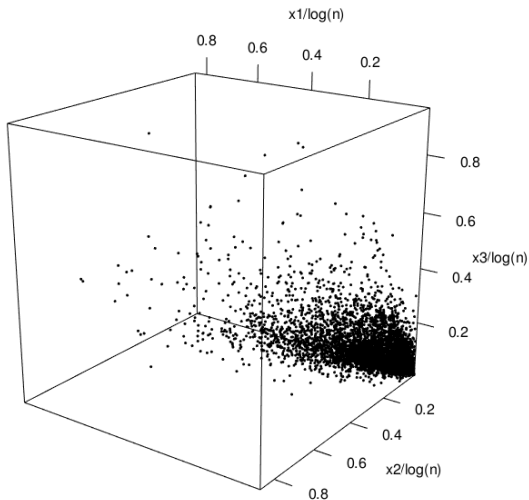


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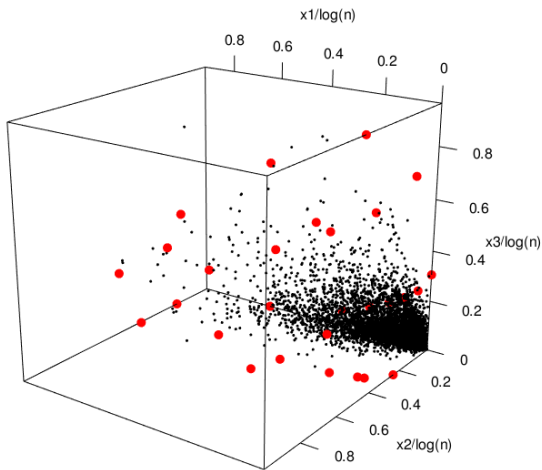


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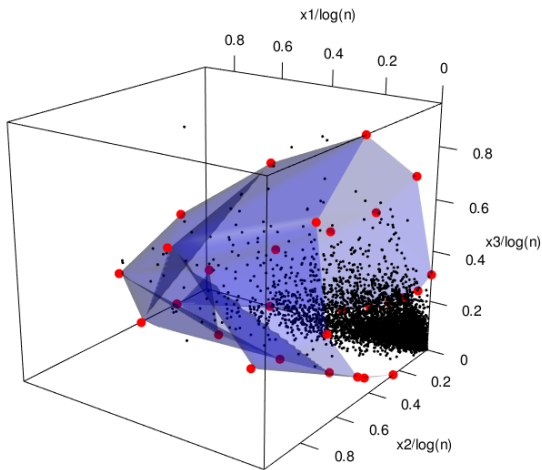


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In d -dimensions...

- ▶ Define a set of N reference angles $\mathbf{w}^{*1}, \dots, \mathbf{w}^{*N} \in \mathcal{S}_{d-1}$.
- ▶ Results in M regions: $\Delta^{(1)}, \Delta^{(2)}, \dots, \Delta^{(M)}$.



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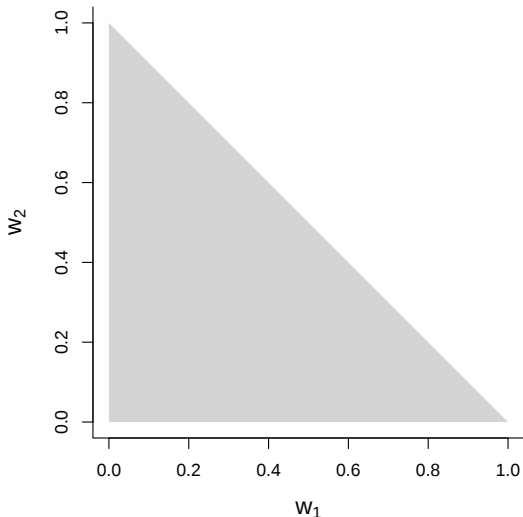
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- ▶ Results in M regions: $\Delta^{(1)}, \Delta^{(2)}, \dots, \Delta^{(M)}$.
- ▶ At a point $\mathbf{x} \in \mathbb{R}^d$, the gauge function value is given by

$$g_{\text{PWL}}(\mathbf{x}; \boldsymbol{\theta}) = \sum_{k=1}^M \mathbf{1}_{\Delta^{(k)}}(\mathbf{x}/\|\mathbf{x}\|) \frac{\mathbf{n}^{(k)\top} \mathbf{x}}{\mathbf{n}^{(k)\top} \boldsymbol{\theta}_1^{(k)} \mathbf{w}^{*(k),1}}$$



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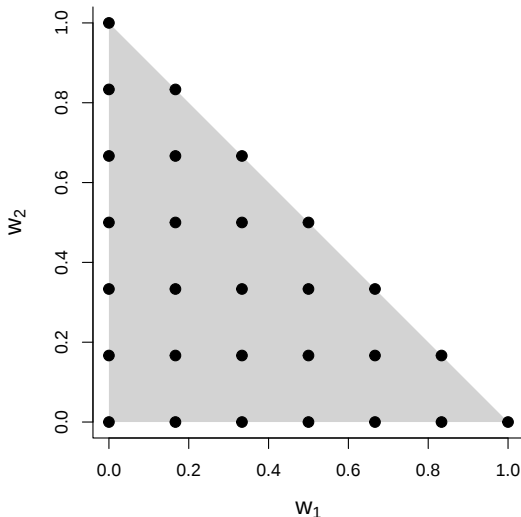
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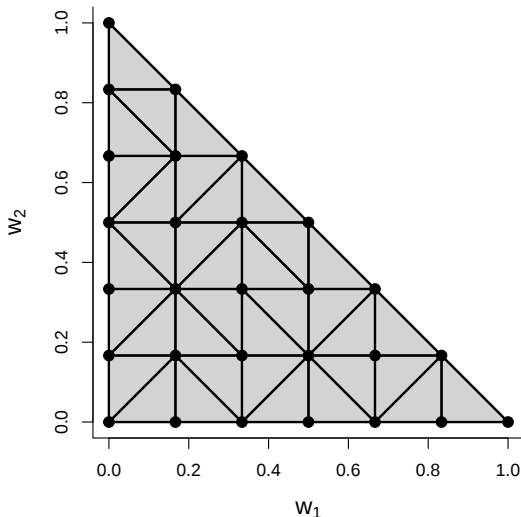
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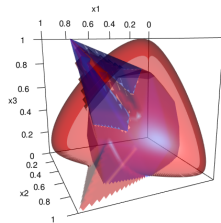
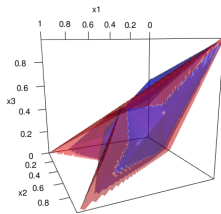
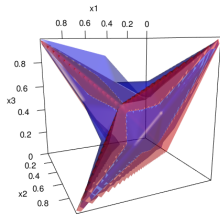
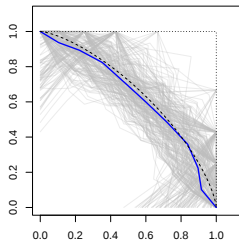
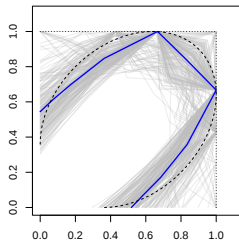
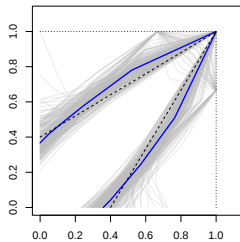
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Geometric approach: extrapolation

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- ▶ Compute $\Pr(\mathbf{X} \in B | R > r_q(\mathbf{W}))$ via sampling:
 1. sample $\mathbf{w}_1, \dots, \mathbf{w}_N$ from $\mathbf{W} \mid \{R > r_q(\mathbf{W})\}$.
 2. sample r_i from $\text{truncGamma}(\hat{\alpha}, g(\mathbf{w}_i; \hat{\theta}))$ for $i = 1, \dots, N$
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- ▶ Compute $\Pr(R > r_q(\mathbf{W}))$ empirically.
- ▶ **Bonus!** Can extrapolate far into the tails using

$$\Pr(\mathbf{X} \in B) = \Pr(\mathbf{X} \in B | R > kr_q(\mathbf{W})) \Pr(R > kr_q(\mathbf{W}))$$

for $k > 1$.



Angular model

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$$f_{\mathbf{W} \mid \{R > r_q(\mathbf{W})\}}(\mathbf{w}; \theta) = \frac{g(\mathbf{w}; \theta)^{-d}}{d\text{vol}(\{\mathbf{x} : g(\mathbf{x}; \theta) \leq 1\})}$$



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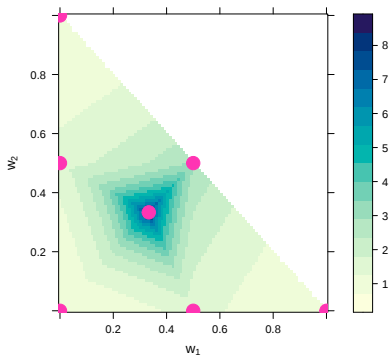
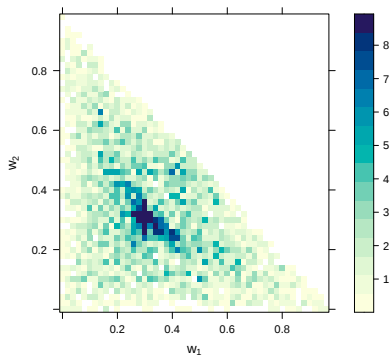
- ▶ Campbell and Wadsworth (2024):
 - ▶ $\text{vol}(\{\mathbf{x} : g_{\text{PWL}}(\mathbf{x}; \boldsymbol{\theta}) \leq 1\})$ has a closed-form expression.
 - ▶ samples are drawn from $f_{\mathbf{W} \mid \{R > r_q(\mathbf{W})\}}$ using MCMC.



Angular model

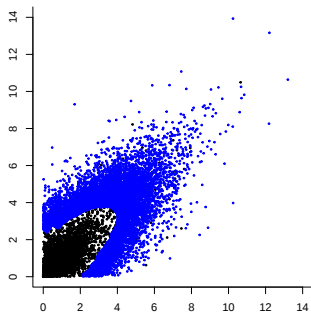
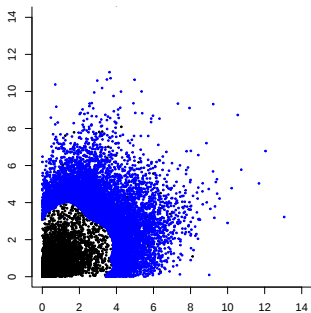
Fitting the angular density

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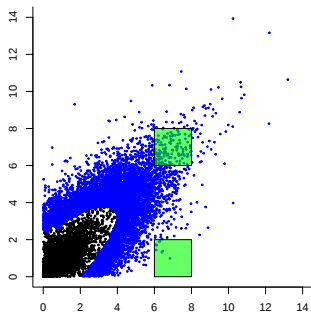
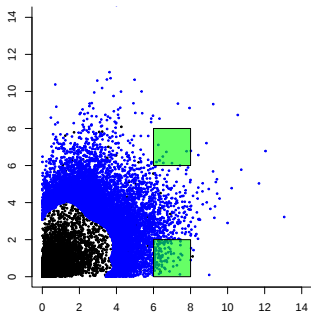


Geometric approach: extrapolation



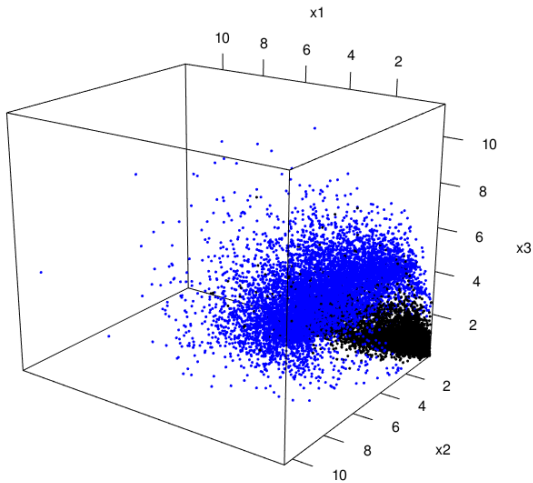


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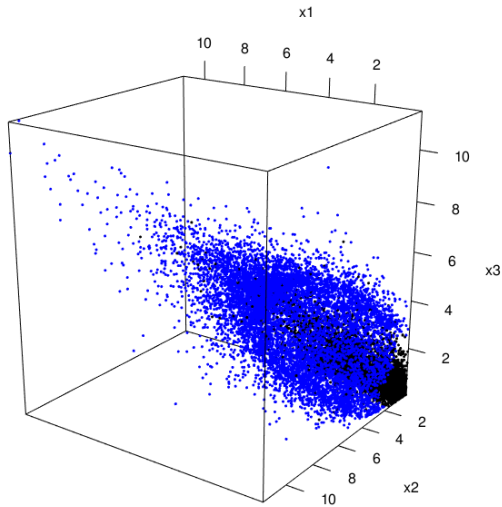


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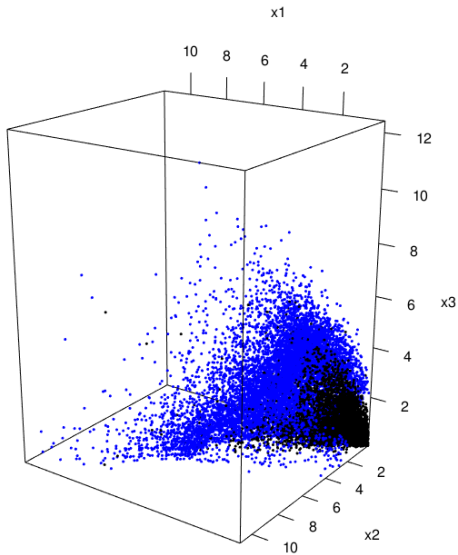


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Application to air pollution measurements ($d = 3$)

- ▶ North Kensington site, London, UK
- ▶ carbon monoxide (CO, mg/m^3), nitrogen dioxide (NO_2 , $\mu\text{g}/\text{m}^3$), and particles with a diameter of $10\ \mu\text{m}$ or less (PM10, mg/m^3).
- ▶ $n = 5,584$ daily maximum measurements, October–April. 1996–2024.

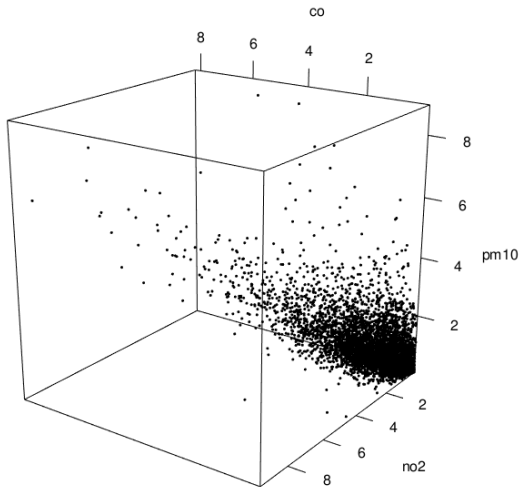


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- ▶ $n = 5,584$ daily maximum measurements, October–April. 1996–2024.
- ▶ Using methods from Simpson et al. (2020), evidence that
 - ▶ $\{\text{PM10}\}$ is large when $\{\text{CO}, \text{NO}_2\}$ are small
 - ▶ $\{\text{CO}, \text{NO}_2, \text{PM10}\}$ grow large together.

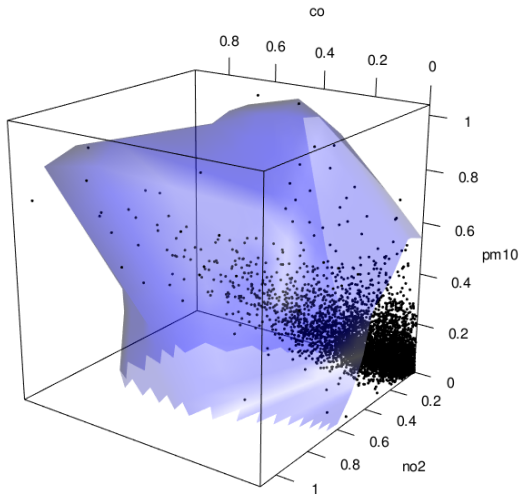


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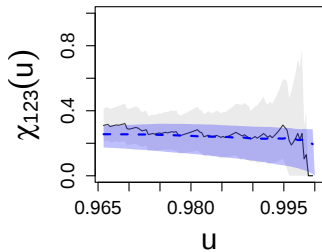
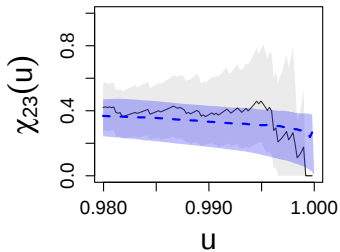
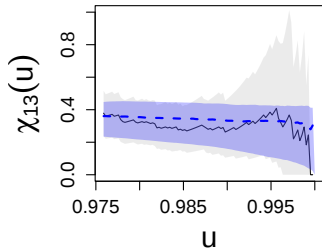
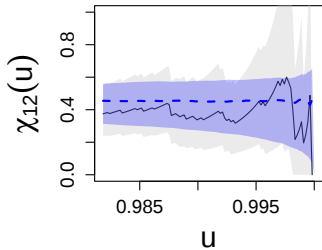




Application to air pollution measurements ($d = 3$)



 $\chi_C(u)$ estimates, $C \subseteq \{1, 2, 3\}$

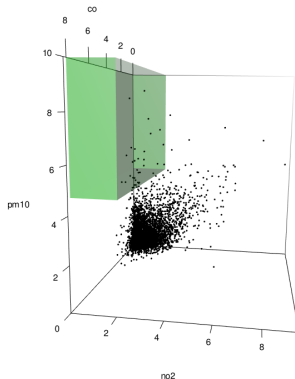
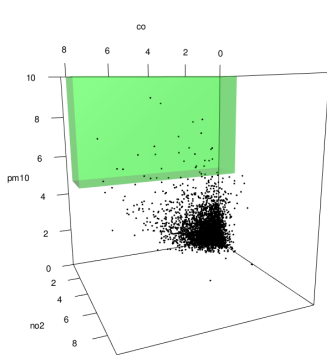


An alternate tail probability

For $d = 3$ pollution data, consider

$$\psi_{\{3\}}(u; \delta_1, \delta_2) = \Pr[F_E(X_1) < \delta_1, F_E(X_2) < \delta_2, F_E(X_3) > u]$$

for $\delta_1, \delta_2 \in [0, 1)$, $u > u_0$, $u_0 \in [0, 1)$ close to 1.

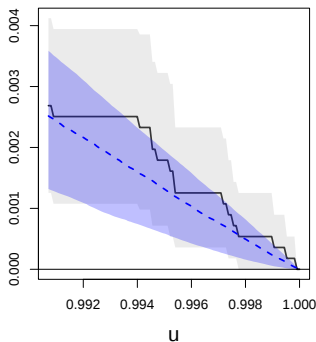


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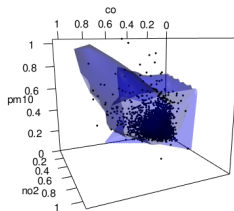
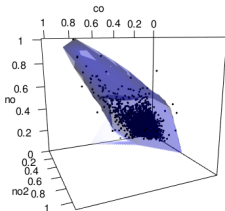
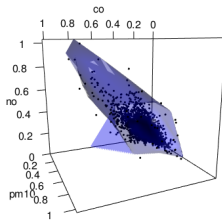
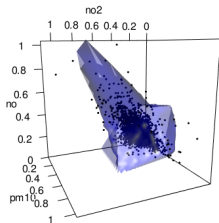
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Application to air pollution measurements ($d = 4$)



Thank you!



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